# Learning, Probabilities and Causality 

Chapter I - Introduction and Conditional Independence

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(2) Introduction and Motivation

3 The multivariate Gaussian Distribution
4) Latent Variables and Conditional Independence

- Conditional Independence
- D-separation
- Markovian dependencies


## Course Organisation

## Course Content

Learning Probabilities and Causality is structured in two parts.
(1) Learning for probabilistic models given a causality graph (Thomas \& Xavi)
(2) Methods for inferring this graph from data (Emilie \& Eric)


## Course Content - Part 1

Probabilistic learning with latent variables:
(1) Basics of probabilistic models: conditional independence
(2) Model-based clustering and Gaussian mixture models
(3) Sequential data and hidden Markov models
(4) Probabilistic principal component analysis
(5) Linear dynamical systems
(6) Approximate variational inference

## LPC Part 1 - Instructors



Research Scientist [website] @thomashueber thomas.hueber@gipsa-lab.fr Leader of CRISSP (cognitive robotics, interactive systems, speech processing), GIPSA-Lab, CNRS

Multimodal speech processing, machine learning, interactive systems

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Audio-visual perception, probabilistic and deep learning, human-robot interaction

## Grading Rules

- Lab work (LW), mid-term exam (ME) final exam (FE).
- Grade $=(\mathrm{LW}+\mathrm{ME}+\mathrm{FE}) / 3$.
- LW = average of all lab works.


## Support material:

https://chamilo.grenoble-inp.fr/courses/ENSIMAGWMM9AM46.

## Calendar

| When? | Who? | Where? | Comment |
| ---: | :---: | :---: | :---: |
| 30-Sep | Xavi | D211 |  |
| 7-Oct | Thomas | D207 |  |
| 14-Oct | Thomas | D207 \& E303 | Lab (15h30-17h) |
| 21-Oct | Xavi | D111 |  |
| 28-Oct | Xavi | H105 |  |
| 18-Nov | Xavi | D111/E303 | Lab (15h30-17h) |
| 25-Nov | Emilie/Eric | H105 | Mid-term Exam (14h-15h30) |
| 2-Dec | Emilie/Charles | H105 |  |
| 9-Dec | Emilie | E201 | Lab (14h-17h) |
| 16-Dec | Eric/Charles | H105 |  |
| 6-Jan | Charles/Emilie | H105 |  |
| 13 Jan | Charles/Eric | H105/E200 | Lab (15h30-17h) |

## References

There is a lot of bibliography on probabilistic graphical models.
I strongly suggest the following book:

Pattern Recognition and Machine Learning, from Christopher M. Bishop (Springer)
The concepts discussed in FPDM correspond to different parts of Ch. 2, 8, 9, 10, 12, 13.
You will not find the part on variational autoencoders (last chapter).

## Introduction and Motivation

## What is probabilistic data mining?

- Probabilistic means we model our data using probabilities.
- For example in classification, we aim to estimate the posterior probability: $P(c \mid x)$ for every possible class $c$.
- Probabilistic generative models \& Bayes rule:

$$
p(x, c)=p(x \mid c) p(c) \Rightarrow p(c \mid x)=\frac{p(x \mid c) p(c)}{\sum_{k} p(x \mid k) p(k)}
$$

- What are all these " $p$ "? What do they mean?


## Probabilities

For discrete variables (i.e. measurable events are discrete):

$$
p(c)=P(C=c)
$$

$\rightarrow$ the probability of the random variable $C$ to value event $c$.

For continuous variables (i.e. measurable events are continuous):

$$
p(x)=f_{X}(x) \quad \text { and } \quad p(\mathcal{X})=P(x \in \mathcal{X})=\int_{\mathcal{X}} f_{X}(x) \mathrm{d} x
$$

$f_{X}$ is the probability density function. Remember $P(\{x\})=0$.

## Why probabilistic data mining?

- Infer hidden variables / exploit partly missing data
- Example: clustering, image segmentation
- Incorporate particular requirements in clustering
- Model complex data (on grids, graphs, temporal, ...)
- Simulate phenomena (speech synthesis), make predictions (regime switching in time series)


Segmentation of time series with respect to the variance

## Example (I): Clustering

- Data: points $\left(x_{j}\right)_{j=1, \ldots, n}$ in $\mathbb{R}^{d}$.
- Aim: find (\& predict) clusters.
- Model-based approach: let $z_{j}$ be the (unknown) cluster of $x_{j}$.
$z_{i}=z_{j} \Rightarrow x_{i}$ and $x_{j}$ should have the same (conditional) distribution.



## Example (II): Dimensionality reduction

- Raw data are high-dimensional descriptors.
- Difficult to mine patterns/visualize.
- Projection on the directions of maximum variance.
- What if for most or even every point $x_{j}$, some coordinates are missing?
- Probabilistic (i.e., model-based) PCA relies on a generative model to exploit partially observed / unknown data.



## Example (III): Analysis of sequential data

- Special case of clustering with temporal dependencies
- Piecewise statistically invariant features with Markovian jumps
- Markovian: it depends only on a few close neighbors.




## The multivariate Gaussian Distribution

## 1D Gaussians

Let's recall the definition of the univariate Gaussian distribution, for $x \in \mathbb{R}$ :

$$
p(x)=\mathcal{N}(x ; \mu, \nu)=\frac{1}{\sqrt{2 \pi \nu}} \exp \left(-\frac{(x-\mu)^{2}}{2 \nu}\right)
$$

- $\mu=\mathbb{E}_{\mathcal{N}(x ; \mu, \nu)}\{x\} \in \mathbb{R}$ is the mean.
- $\nu=\mathbb{E}_{\mathcal{N}(x ; \mu, \nu)}\left\{(x-\mu)^{2}\right\} \in \mathbb{R}^{+}$is the variance.


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Remark 1.1: We will often use the expectation of a function $f$ of a random variable $x$ w.r.t. the probability density function $p(x)$, and denote it by:

$$
\mathbb{E}_{p(x)}\{f(x)\}=\int_{\mathcal{X}} f(x) p(x) \mathrm{d} x
$$

where $\mathcal{X}$ is the domain of the random variable $x$.

## 1D Gaussians: ML estimators

And of its ML estimators for a set of $N$ samples $X=\left\{x_{1}, \ldots, x_{N}\right\}$ :

$$
\mathcal{L}(\mu, \nu \mid X)=\sum_{n=1}^{N} \log \mathcal{N}\left(x_{n} ; \mu, \nu\right)
$$

Exercise 1.1: Prove that the maximum likelihood estimators are:

$$
\mu^{*}=\frac{1}{N} \sum_{n=1}^{N} x_{n} \quad \nu^{*}=\frac{1}{N} \sum_{n=1}^{N}\left(x_{n}-\mu^{*}\right)^{2} .
$$

Hint: Compute $\frac{\partial \mathcal{L}}{\partial \mu}$ and $\frac{\partial \mathcal{L}}{\partial \nu}$ knwoing that $\log \mathcal{N}(x ; \mu, \nu)=-\frac{1}{2}\left(\log (2 \pi \nu)+\frac{(x-\mu)^{2}}{\nu}\right)$.

## Proof

$$
\begin{aligned}
\mathcal{L}(\mu, \nu \mid X) & =\sum_{n=1}^{N} \log \mathcal{N}\left(x_{n} ; \mu, \nu\right) \\
& =-\frac{1}{2} \sum_{n=1}^{N} \log (2 \pi \nu)+\frac{\left(x_{n}-\mu\right)^{2}}{\nu} \\
& =-\frac{1}{2}\left(N \log (2 \pi \nu)+\frac{1}{\nu} \sum_{n=1}^{N}\left(x_{n}-\mu\right)^{2}\right)
\end{aligned}
$$

And hence:

$$
\frac{\partial \mathcal{L}}{\partial \mu}=\frac{1}{\nu} \sum_{n=1}^{N}\left(x_{n}-\mu\right) \quad \frac{\partial \mathcal{L}}{\partial \nu}=-\frac{N}{2 \nu}+\frac{1}{2 \nu^{2}} \sum_{n=1}^{N}\left(x_{n}-\mu\right)^{2}
$$

By setting the derivatives to 0 , we obtain the sought result.

## Multivariate Gaussian distribution

Trivial extension: consider each dimension independently.

$$
p(\mathbf{x})=\mathcal{N}(\mathbf{x} ; \boldsymbol{\mu}, \boldsymbol{\nu})=\prod_{d=1}^{D} \frac{1}{\sqrt{2 \pi \nu_{d}}} \exp \left(-\frac{\left(x_{d}-\mu_{d}\right)^{2}}{2 \nu_{d}}\right)
$$

with $\mathbf{x}=\left(x_{1}, \ldots, x_{D}\right) \in \mathbb{R}^{d}, \boldsymbol{\mu}=\left(\mu_{1}, \ldots, \mu_{D}\right) \in \mathbb{R}^{d}$ and $\boldsymbol{\nu}=\left(\nu_{1}, \ldots, \nu_{D}\right) \in \mathbb{R}^{+}$.


Exercise 1.2: Derive the maximum likelihood estimators for $\mu$ and $\nu$.

## Multivariate Gaussian distribution (II)

By defining:

$$
\boldsymbol{\Sigma}=\operatorname{diag}(\boldsymbol{\nu})=\left(\begin{array}{cccc}
\nu_{1} & 0 & \ldots & 0 \\
0 & \nu_{2} & \ldots & 0 \\
\vdots & \vdots & \ddots & \vdots \\
0 & 0 & \ldots & \nu_{D}
\end{array}\right)
$$

the density rewrites as:

$$
p(\mathbf{x})=\mathcal{N}(\mathbf{x} ; \boldsymbol{\mu}, \boldsymbol{\Sigma})=\frac{1}{\sqrt{|2 \pi \boldsymbol{\Sigma}|}} \exp \left(-\frac{1}{2}(\boldsymbol{x}-\boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1}(\boldsymbol{x}-\boldsymbol{\mu})\right)
$$

Exercise 1.3: Prove it!

## Symmetric and positive definite matrices

Question: does it work for any matrix $\boldsymbol{\Sigma}$ ?
Only for symmetric and positive definite (s.p.d.) matrices.
Remark 1.2: A $D \times D$ symmetric matrix $\boldsymbol{\Sigma}$ is positive definite if and only if $\boldsymbol{v}^{\top} \boldsymbol{\Sigma} \boldsymbol{v}>0, \forall \boldsymbol{v} \neq \mathbf{0}$.

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For the Gaussian distribution, this is intuitive, since the variance should be strictly positive in any direction:


## Symmetric and positive definite matrices (II)

Let $\boldsymbol{\Sigma}$ be a s.p.d. matrix:

- All eigenvalues of $\boldsymbol{\Sigma}$ are ...?


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- How do we know that $\boldsymbol{\Sigma}^{-1}$ exists?


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- How do we know that $\boldsymbol{\Sigma}^{-1}$ exists? The determinant is the product of eigenvalues.
- We can write $\boldsymbol{\Sigma}=\boldsymbol{U} \boldsymbol{\Lambda} \boldsymbol{U}^{\top}$ with $\boldsymbol{\Lambda}$ diagonal and $\boldsymbol{U}$ orthogonal. Why?


## Symmetric and positive definite matrices (II)

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- So $\boldsymbol{\Lambda}$ contains eigenvalues and $\boldsymbol{U}$ contains eigenvectors (as columns).


## Symmetric and positive definite matrices (II)

Let $\boldsymbol{\Sigma}$ be a s.p.d. matrix:

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- We can write $\boldsymbol{\Sigma}=\boldsymbol{U} \boldsymbol{\Lambda} \boldsymbol{U}^{\top}$ with $\boldsymbol{\Lambda}$ diagonal and $\boldsymbol{U}$ orthogonal. Why?
- So $\boldsymbol{\Lambda}$ contains eigenvalues and $\boldsymbol{U}$ contains eigenvectors (as columns).
- Then, $\boldsymbol{\Sigma}^{-1}=\boldsymbol{U} \boldsymbol{\Lambda}^{-1} \boldsymbol{U}^{\top}$.

Exercise 1.4: Prove that the inverse of a (symmetric) positive definite matrix always exists.

## Defining multivariate Gaussians

Remark 1.3: Given a vector $\mu \in \mathbb{R}^{D}$ and a s.p.d. matrix $\boldsymbol{\Sigma} \in \mathbb{R}^{D \times D}$, we can define the multivariate Gaussian distribution as:

$$
\mathcal{N}(\mathbf{x} ; \boldsymbol{\mu}, \boldsymbol{\Sigma})=\frac{1}{\sqrt{|2 \pi \boldsymbol{\Sigma}|}} \exp \left(-\frac{1}{2}(\boldsymbol{x}-\boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1}(\boldsymbol{x}-\boldsymbol{\mu})\right)
$$

$\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$ are usually referred to as the mean vector and the covariance matrix, and they are defined as:

$$
\boldsymbol{\mu}=\mathbb{E}_{\mathcal{N}(\boldsymbol{x} ; \boldsymbol{\mu}, \boldsymbol{\Sigma})}\{\mathbf{x}\} \quad \boldsymbol{\Sigma}=\mathbb{E}_{\mathcal{N}(\boldsymbol{x} ; \boldsymbol{\mu}, \boldsymbol{\Sigma})}\left\{(\mathbf{x}-\boldsymbol{\mu})(\mathbf{x}-\boldsymbol{\mu})^{\top}\right\} .
$$

Exercise 1.5: Prove that the normalisation constant of a multivariate Gaussian distribution with covariance matrix $\boldsymbol{\Sigma}$ is $\sqrt{|2 \pi \boldsymbol{\Sigma}|}$.

Jupyter Notebook!!!

## Standard Gaussian and Affine Transforms

Remark 1.4: The standard multivariate Gaussian is defined as the zeromean and unit-variance Gaussian distribution:

$$
\mathcal{N}(\mathbf{z} ; \mathbf{0}, \mathbf{I})=\frac{1}{(2 \pi)^{D / 2}} \exp \left(-\frac{1}{2}\|\mathbf{z}\|^{2}\right)
$$

Exercise 1.6: Let us consider the case where $\mathbf{z}$ follows a standard multivariate Gaussian distribution, and we define $\mathbf{x}=\mathbf{A z}+\boldsymbol{\mu}$ with $\mathbf{A} \in \mathbb{R}^{D \times D}$ being an invertible matrix $(|\mathbf{A}| \neq 0)$. Prove that:

$$
p(\mathbf{x})=\mathcal{N}(\mathbf{x} ; \boldsymbol{\mu}, \boldsymbol{\Sigma}), \quad \text { with } \quad \boldsymbol{\Sigma}=\mathbf{A A}^{\top} .
$$

## More on Multivariate Gaussians

What are the level curves of the Gaussian p.d.f.?

$$
\mathcal{C}_{\lambda}=\{\boldsymbol{x} \mid \mathcal{N}(\boldsymbol{x} ; \boldsymbol{\mu}, \boldsymbol{\Sigma})=\lambda\}
$$

## More on Multivariate Gaussians

What are the level curves of the Gaussian p.d.f.?

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- Empty set for $\lambda<0$.
- $\mathcal{C}_{0}=\{\boldsymbol{\mu}\}$.
- $\mathcal{C}_{\lambda}$ ?


## More on Multivariate Gaussians

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$$

- Empty set for $\lambda<0$.
- $\mathcal{C}_{0}=\{\boldsymbol{\mu}\}$.
- $\mathcal{C}_{\lambda}$ ? An ellipsoid with center $\mu$ with axis given by the columns of $\boldsymbol{U}$ and axis length given by the elements in $\boldsymbol{\Lambda}$, where $\boldsymbol{\Sigma}=\boldsymbol{U} \boldsymbol{\Lambda} \boldsymbol{U}^{\top}$.



## ML for multivariate Gaussians

Exercise 1.7: Prove that the ML estimators of the multivariate Gaussian are:

$$
\boldsymbol{\mu}^{*}=\frac{1}{N} \sum_{n=1}^{N} \boldsymbol{x}_{n} \quad \boldsymbol{\Sigma}^{*}=\frac{1}{N} \sum_{n=1}^{N}\left(\boldsymbol{x}_{n}-\boldsymbol{\mu}^{*}\right)\left(\boldsymbol{x}_{n}-\boldsymbol{\mu}^{*}\right)^{\top} .
$$

## ML for multivariate Gaussians

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$$

You will need to take derivatives w.r.t. matrices. Let $\boldsymbol{M}$ be a matrix, and $f(\boldsymbol{M})$ and function of that matrix (e.g. trace, ...). One can consider $\frac{\partial f}{\partial \boldsymbol{M}}$.

Examples of matrix derivative formulae useful to derive the ML estimate of multivariate Gaussians:

$$
\begin{gathered}
\frac{\partial \operatorname{Tr}(\boldsymbol{M} \boldsymbol{A})}{\partial \boldsymbol{M}}=\boldsymbol{A}^{\top} \quad \frac{\partial \operatorname{Tr}\left(\boldsymbol{B}^{\top} \boldsymbol{M}^{\top} \boldsymbol{C M B}\right)}{\partial \boldsymbol{M}}=\boldsymbol{C}^{\top} \boldsymbol{\boldsymbol { M B B } ^ { \top } + \boldsymbol { C M B B } ^ { \top }} \\
\frac{\partial \log |\boldsymbol{M}|}{\partial \boldsymbol{M}}=\left(\boldsymbol{M}^{-1}\right)^{\top} \quad[\operatorname{Tr}(\boldsymbol{A B C})=\operatorname{Tr}(\boldsymbol{B C A})=\operatorname{Tr}(\boldsymbol{C A B})]
\end{gathered}
$$

## Gaussian Completion: Shape is All You Need

Remark 1.5: Developing multivariate Gaussian distribution, we observe that only two terms depend on $\mathbf{x}$ (quadratic and linear):

$$
\mathcal{N}(\boldsymbol{x} ; \boldsymbol{\mu}, \boldsymbol{\Sigma}) \stackrel{\mathbf{x}}{\propto} \exp \left(-\frac{1}{2} \boldsymbol{x}^{\top} \boldsymbol{\Sigma}^{-1} \boldsymbol{x}+\mathbf{x}^{\top} \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}\right)
$$

$(\stackrel{x}{\propto}$ means that is proportional up to a constant that does NOT depend on $\mathbf{x}$ )

## Gaussian Completion: Shape is All You Need

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$$

$(\stackrel{x}{\propto}$ means that is proportional up to a constant that does NOT depend on $\mathbf{x}$ )
Exercise 1.8: Prove that given a s.p.d. matrix $\Omega$ and a vector $\mathbf{m}$ :

$$
p(\mathbf{x}) \stackrel{\mathbf{x}}{\propto} \exp \left(-\frac{1}{2} \mathbf{x}^{\top} \boldsymbol{\Omega} \mathbf{x}+\mathbf{x}^{\top} \mathbf{m}\right) \Rightarrow p(\mathbf{x})=\mathcal{N}(\mathbf{x} ; \boldsymbol{\mu}, \boldsymbol{\Sigma})
$$

with:

$$
\boldsymbol{\Sigma}=\mathbf{\Omega}^{-1} \quad \boldsymbol{\mu}=\boldsymbol{\Sigma} \mathbf{m}=\mathbf{\Omega}^{-1} \mathbf{m}
$$

More on multivariate Gaussians in Chapter 4.

## Latent Variables and Conditional Independence

## What is a model?

What does it mean to model the relationship between two variables?

- We choose the nature of $\boldsymbol{z} \& \boldsymbol{x}$ : cont./discrete, bounded, ...
- We choose the dependencies, i.e. $p(\boldsymbol{x}, \boldsymbol{z})=p(\boldsymbol{x} \mid \boldsymbol{z}) p(\boldsymbol{z})$.
- We choose the prior distribution $p(\boldsymbol{z})$.
- We choose the likelihood distribution $p(\boldsymbol{x} \mid \boldsymbol{z})$.


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Remark 1.6: There is an important difference between observed and latent or hidden variables. Observed variables are measured, and latent variables are quantities that cannot be measured directly.

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- We choose the likelihood distribution $p(\mathbf{x} \mid \boldsymbol{z})$.

Remark 1.7: In models with latent variables, we study the marginal distribution of $\mathbf{x}$ (left) and the posterior distribution of $\mathbf{z}$ given $\mathbf{x}$ (right):

$$
p(\mathbf{x})=\int_{\mathcal{Z}} p(\mathbf{x} \mid \mathbf{z}) p(\mathbf{z}) \mathrm{d} \mathbf{z} \quad p(\mathbf{z} \mid \mathbf{x})=\frac{p(\mathbf{x} \mid \mathbf{z}) p(\mathbf{z})}{p(\mathbf{x})}
$$

## Example: Gaussian mixture model

- The nature: $z$ is discrete \& bounded, $x$ is 1D \& continuous.
- The dependencies: $p(x, z)=p(x \mid z) p(z)$.


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- The nature: $z$ is discrete \& bounded, $x$ is 1D \& continuous.
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- The distribution $p(z), z \in\{1, \ldots, K\}$ is categorical:

$$
p(z=k)=\pi_{k}, \quad \pi_{k} \geq 0, \sum_{k=1}^{K} \pi_{k}=1
$$

## Example: Gaussian mixture model

- The nature: $z$ is discrete \& bounded, $x$ is 1D \& continuous.
- The dependencies: $p(x, z)=p(x \mid z) p(z)$.
- The distribution $p(z), z \in\{1, \ldots, K\}$ is categorical:

$$
p(z=k)=\pi_{k}, \quad \pi_{k} \geq 0, \sum_{k=1}^{K} \pi_{k}=1
$$

- The distribution $p(x \mid z)$ is Gaussian:

$$
p(x \mid z=k)=\mathcal{N}\left(x ; \mu_{k}, \nu_{k}\right)=\frac{1}{\sqrt{2 \pi \nu_{k}}} \exp \left(-\frac{\left(x-\mu_{k}\right)^{2}}{2 \nu_{k}}\right)
$$

with $\mu_{k} \in \mathbb{R}$ and $\nu_{k}>0, \forall k$.

## Likelihood and GMM posteriors

Exercise 1.9: Prove that the GMM marginal writes:

$$
p(x)=\sum_{k=1}^{K} \pi_{k} \frac{1}{\sqrt{2 \pi \nu_{k}}} \exp \left(-\frac{\left(x-\mu_{k}\right)^{2}}{2 \nu_{k}}\right) .
$$

Exercise 1.10: Prove the GMM posterior writes:

$$
p(z=k \mid x)=\frac{\pi_{k} \frac{1}{\sqrt{2 \pi \nu_{k}}} \exp \left(-\frac{\left(x-\mu_{k}\right)^{2}}{2 \nu_{k}}\right)}{\sum_{m=1}^{K} \pi_{m} \frac{1}{\sqrt{2 \pi \nu_{m}}} \exp \left(-\frac{\left(x-\mu_{m}\right)^{2}}{2 \nu_{m}}\right)}
$$

Hint: Just write down what things are.
More on this on the Chapter 2.

# Latent Variables and Conditional Independence 

 Conditional Independence
## 3-variable models: taxonomy



Full all dependencies are set:

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$

## 3-variable models: taxonomy



Full all dependencies are set:

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$



Two-kids $Y-Z$ dependency missing:

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$

## 3-variable models: taxonomy



Full all dependencies are set:

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$



Two-kids $Y$ - $Z$ dependency missing:

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$

(X) Two-parents $X-Y$ dependency missing:

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$

## 3-variable models: taxonomy



Full all dependencies are set:

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$


(Y) Two-kids $Y-Z$ dependency missing:

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$

(X) Two-parents $X-Y$ dependency missing:

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$



Cascaded $X-Z$ dependency missing:
(Z)

$$
p(x, y, z)=p(z \mid x, y) p(y \mid x) p(x)
$$

## 3-variable models: Two-kids



Two-kids $p(x, y, z)=p(z \mid x) p(y \mid x) p(x)$.

Exercise 1.11: Prove that in the Two-kids model:

$$
p(y \mid z) \neq p(y) \quad \text { and } \quad p(y \mid z, x)=p(y \mid x)
$$

- The first statement is equivalent to say that $y$ and $z$ are not independent.
- The second statement, says that $y$ and $z$ are conditionally independent w.r.t. $x$.


## 3-variable models: Two-parents



Two-parents $p(x, y, z)=p(z \mid x, y) p(x) p(y)$.

Exercise 1.12: Prove that in the Two-parents model:

$$
\begin{equation*}
p(y \mid x)=p(x) \quad \text { and } \quad p(x \mid y, z) \neq p(x \mid z) \tag{1}
\end{equation*}
$$

We are in the opposite case:

- The first statement says that $y$ and $x$ are independent.
- The second statement says that $y$ and $x$ are conditionally dependent w.r.t. z.


## 3-variable models: Cascaded



Cascaded $p(x, y, z)=p(z \mid y) p(y \mid x) p(x)$.

Exercise 1.13: Prove that in the Cascaded model:

$$
\begin{equation*}
p(x, z) \neq p(x) p(z) \quad \text { and } \quad p(x, z \mid y)=p(x \mid y) p(z \mid y) \tag{2}
\end{equation*}
$$

In this case, we obtain similar results than with the Two-kids model.

## 3-variable models: Cascaded



Cascaded $p(x, y, z)=p(z \mid y) p(y \mid x) p(x)$.

Exercise 1.13: Prove that in the Cascaded model:

$$
\begin{equation*}
p(x, z) \neq p(x) p(z) \quad \text { and } \quad p(x, z \mid y)=p(x \mid y) p(z \mid y) \tag{2}
\end{equation*}
$$

In this case, we obtain similar results than with the Two-kids model.

Remark 1.8: At this point it should be clear that independence and conditional independence are two very different properties of random variables.

## Conditional Independence: Definition

Remark 1.9: Let $x, y$, and $z$ be random variables, we say that $x$ and $y$ are conditionally independent given $z$, and write $x \Perp y \mid z$, iff one of the following equivalent expressions holds:

- $p(x, y \mid z)=p(x \mid z) p(y \mid z)$
- $p(x \mid y, z)=p(x \mid z)$
- $p(y \mid x, z)=p(y \mid z)$

Latent Variables and Conditional Independence

## D-separation

## Motivation

Let's consider the following variables and dependencies.


Is $P \Perp V \mid T$ ? How would you do it? Is the previous strategy scalable?

## Basics

Let us recall the 3-var models:
$X \longrightarrow Y$ Two kids The path from $Z$ to $Y$ is called "tail-to-tail."

$$
p(z, y \mid x)=p(z \mid x) p(y \mid x)
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## Path blocking



## Path blocking



The purple node "blocks" the path in two-kids/tail-to-tail \& cascaded/head-to-tail $\rightarrow$ conditional independence.

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The purple node "blocks" the path in two-kids/tail-to-tail \& cascaded/head-to-tail $\rightarrow$ conditional independence.

The purple node "unblocks" the path in two-parents/head-to-head $\rightarrow$ conditional dependence.

## Path blocking



The purple node "blocks" the path in two-kids/tail-to-tail \& cascaded/head-to-tail $\rightarrow$ conditional independence.

The purple node "unblocks" the path in two-parents/head-to-head $\rightarrow$ conditional dependence. In the two-parents, $Z$ or any descendant of $Z$ will unblock the path.

## Path blocking (revisited)

(Let me change the variable names)


Two-parents
head-to-head


## Path blocking (revisited)

(Let me change the variable names)


Tail-to-tail \& head-to-tail $\rightarrow A \Perp B \mid C$.

## Path blocking (revisited)

(Let me change the variable names)


Tail-to-tail \& head-to-tail $\rightarrow A \Perp B \mid C$.
Head-to-head $\rightarrow A \not \perp B \mid C$ or any descendant of $C$.

## Path blocking (revisited)

(Let me change the variable names)


Tail-to-tail \& head-to-tail $\rightarrow A \Perp B \mid C$.
Head-to-head $\rightarrow A \not \Perp B \mid C$ or any descendant of $C$.
$\Rightarrow$ If $A \Perp B \mid C$, nodes within tail-to-tail or head-to-tail can be in $C$ and nodes within head-to-head or any of their descendents must not be in C.

## Path blocking (definition)

Remark 1.10: Let $A, B$ and $C$ be three non-intersecting sets of nodes of directed acyclic graph. A path from $A$ to $B$ is said to be blocked by $C$ if it includes a node that either:

- the path meets tail-to-tail or head-to-tail at the node and the node is in C;
- the path meets head-to-head at the node and neither the node nor any of its descendants are in C .


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(Corresponds to Exercise: 1.14.)
- Is the path from $\{x\}$ to $\{v\}$ blocked by $\{u\}$ ?


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## D-separation

Remark 1.11: Let $A, B$ and $C$ be three non-intersecting sets of nodes of a directed acyclic graph. $A$ and $B$ are $\mathbf{D}$-separated by $C$, if all paths from any node from $A$ to $B$ are blocked by $C$.

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Remark 1.12: $A$ and $B$ are $D$-separated by $C$ if and only if $A \Perp B \mid C$.

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Latent Variables and Conditional Independence
Markovian dependencies

## Markov models: introduction

Principle: each variable depends only on its closer neighbours. Examples:


Markov chain.

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Markov chain (top).
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Markov chain (top). Hidden Markov chain (bottom).


Double hidden Markov chain.

## D-separation in Markov models

With the following model:


Is $\left\{X_{t-1}\right\}$ D-separated from $\left\{Y_{t+1}\right\}$ by ...

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- $\left\{X_{t}\right\}$ ? Yes
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- $\left\{X_{t}\right\}$ ? Yes
- $\left\{Y_{t}\right\}$ ? No
- $\left\{X_{t}, Y_{t}\right\}$ ? [1 minute]


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- $\left\{X_{t}, Y_{t}\right\}$ ? [1 minute]


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With the following model:


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- $\left\{X_{t}\right\}$ ? No
- $\left\{Y_{t}\right\}$ ? No
- $\left\{X_{t}, Y_{t}\right\}$ ? Yes


## D-separation in Markov models: summary

Is $\left\{X_{t-1}\right\}$ D-separated from $\left\{Y_{t+1}\right\}$ by (left column) in (top row)?

$\left\{X_{t}\right\}$
$\left\{Y_{t}\right\}$
$\left\{X_{t}, Y_{t}\right\}$
Yes
No
Yes
No
Yes
Yes


No
No
Yes
(Corresponds to Exercise 1.16.)

## Markov blanket (or boundary)

Model example:


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For a given node $K$, what is the minimal set of variables $\mathcal{B}_{K}$ so that:

$$
p(K \mid \text { all except } K)=p\left(K \mid \mathcal{B}_{K}\right) ?
$$

You've got 3 minutes!

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For $L$ ? $\quad \mathcal{B}_{L}=\{F, G, H\}$ because $F, G, H$ are parents of $L$.

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For a given node $K$, what is the minimal set of variables $\mathcal{B}_{K}$ so that:

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$$

For $L$ ? $\quad \mathcal{B}_{L}=\{F, G, H\}$ because $F, G, H$ are parents of $L$.
For $C$ ? $\quad \mathcal{B}_{C}=\{B, D, F\}$ because $B(F, D)$ is parent (children) of $C$.

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Model example:


For a given node $K$, what is the minimal set of variables $\mathcal{B}_{K}$ so that:

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For $L$ ? $\quad \mathcal{B}_{L}=\{F, G, H\}$ because $F, G, H$ are parents of $L$.
For $C$ ? $\quad \mathcal{B}_{C}=\{B, D, F\}$ because $B(F, D)$ is parent (children) of $C$.
For $E$ ? You've got 3 minutes!

## Markov blanket (or boundary)

Model example:


For a given node $K$, what is the minimal set of variables $\mathcal{B}_{K}$ so that:

$$
p(K \mid \text { all except } K)=p\left(K \mid \mathcal{B}_{K}\right) ? \quad \mathcal{B}_{K}=\{G\}
$$

For $L$ ? $\quad \mathcal{B}_{L}=\{F, G, H\}$ because $F, G, H$ are parents of $L$.
For $C$ ? $\quad \mathcal{B}_{C}=\{B, D, F\}$ because $B(F, D)$ is parent (children) of $C$.
For $E$ ? $\quad \mathcal{B}_{E}=\{A, B, I, J\}$ because $A, B(I)$ are parents (children) of $E$ and $J$ is co-parent of $E$.

## Markov blanket: definition

Remark 1.13: The Markov blanket is the minimal set that D-separates a set of nodes from the rest of the graph.


Remark 1.14: Construction of the Markov blanket. Given a directed acyclic graph, and a node $x$ on that graph, the Markov blanket of $x, \mathcal{B}_{x}$ is the set of all parents, children and co-parents of $x$.

